

Joint distribution of new sample rank of bivariate order statistics

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Appendix

Proof of Lemma 1. We consider the limit

$$f_{r,s;n}(x, y) = \lim_{\delta_x, \delta_y \rightarrow 0} \frac{P\{x < X_{r:n} \leq x + \delta_x, y < Y_{s:n} \leq y + \delta_y\}}{\delta_x \delta_y},$$

$$-\infty < x, y < \infty \quad (1)$$

and use the description of compound event $\{x < X_{r:n} \leq x + \delta_x, y < Y_{s:n} \leq y + \delta_y\}$ with the multinomial distribution to obtain:

$$\begin{aligned} f_{r,s;n}(x, y) \delta_x \delta_y &\approx P\{x < X_{r:n} \leq x + \delta_x, y < Y_{s:n} \leq y + \delta_y\} \\ &= P\{r - 1 \text{ of } X's \in (-\infty, x], 1 \text{ of } X's \in (x, x + \delta_x], \\ &\quad n - r \text{ of } X's \in (x + \delta_x, \infty), s - 1 \text{ of } Y's \in (-\infty, y], \\ &\quad 1 \text{ of } Y's \in (y, y + \delta_y], n - s \text{ of } Y's \in (y + \delta_y, \infty)\}. \end{aligned}$$

The compound event $\{x < X_{r:n} \leq x + \delta_x, y < Y_{s:n} \leq y + \delta_y\}$ can be realized by the following configuration:

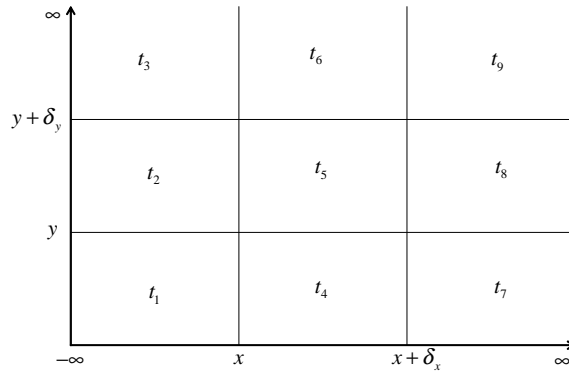


Figure 1. Realization of the compound event

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This means that the total number of observations (X, Y) falls in nine regions denoted by I_i , where $I_i \cap I_j = \emptyset$ and $I_i \cup I_j = \overline{\mathbb{R}}^2$; $\overline{\mathbb{R}} = \mathbb{R} \cup \{-\infty, \infty\}$ for all $i \neq j$, $i, j = 1, 2, \dots, 9$. All possible regions I_1, I_2, \dots, I_9 are defined as follows:

$$\begin{aligned} I_1 &= (-\infty, x] \cap (-\infty, y], & I_6 &= (x, x + \delta_x] \cap (y + \delta_y, \infty), \\ I_2 &= (-\infty, x] \cap (y, y + \delta_y], & I_7 &= (x + \delta_x, \infty) \cap (-\infty, y], \\ I_3 &= (-\infty, x] \cap (y + \delta_y, \infty), & I_8 &= (x + \delta_x, \infty) \cap (y, y + \delta_y], \\ I_4 &= (x, x + \delta_x] \cap (-\infty, y], & I_9 &= (x + \delta_x, \infty) \cap (y + \delta_y, \infty), \\ I_5 &= (x, x + \delta_x] \cap (y, y + \delta_y], \end{aligned}$$

Let t_i observations fall in the region I_i , $i = 1, 2, \dots, 9$. For example, t_1 of (X, Y) are observed in $I_1 = (-\infty, x] \cap (-\infty, y]$ and t_2 of (X, Y) are observed in $I_2 = (-\infty, x] \cap (y, y + \delta_y]$.

Let us denote $C_1 = \{X \leq x, Y \leq y\}$, $C_2 = \{X \leq x, y < Y \leq y + \delta_y\}$, $C_3 = \{X \leq x, Y > y + \delta_y\}$, $C_4 = \{x < X \leq x + \delta_x, Y \leq y\}$, $C_5 = \{x < X \leq x + \delta_x, y < Y \leq y + \delta_y\}$, $C_6 = \{x < X \leq x + \delta_x, Y > y + \delta_y\}$, $C_7 = \{X > x + \delta_x, Y \leq y\}$, $C_8 = \{X > x + \delta_x, y < Y \leq y + \delta_y\}$, $C_9 = \{x < X \leq x + \delta_x, y < Y \leq y + \delta_y\}$.

In the experiment, assume that the outcomes are pairs C_i with probabilities $P(C_i) = p_i$, where $\sum_i p_i = 1$ and $i = 1, 2, \dots, 9$. Then we have

$$\begin{aligned} P(C_1) &= F(x, y), \\ P(C_2) &= F(x, y + \delta_y) - F(x, y), \\ P(C_3) &= F_X(x) - F(x, y + \delta_y), \\ P(C_4) &= F(x + \delta_x, y) - F(x, y), \\ P(C_5) &= F(x + \delta_x, y + \delta_y) - F(x + \delta_x, y) - F(x, y + \delta_y) + F(x, y), \\ P(C_6) &= F_X(x + \delta_x) - F_X(x) - F(x + \delta_x, y + \delta_y) + F(x, y + \delta_y), \\ P(C_7) &= F_Y(y) - F(x + \delta_x, y), \\ P(C_8) &= F_Y(y + \delta_y) - F_Y(y) - F(x + \delta_x, y + \delta_y) + F(x + \delta_x, y), \\ P(C_9) &= 1 - F_X(x + \delta_x) - F_Y(y + \delta_y) + F(x + \delta_x, y + \delta_y). \end{aligned}$$

Let ζ_i be the number of occurrences in which C_i appears out of n repetitions, $i = 1, 2, \dots, 9$. Clearly, the random vector $(\zeta_1, \zeta_2, \dots, \zeta_9)$ is multinomial with pmf

$$\begin{aligned} P\{\zeta_1 = t_1, \zeta_2 = t_2, \dots, \zeta_9 = t_9\} \\ = \frac{n!}{t_1! t_2! \dots t_9!} [P(C_1)]^{t_1} [P(C_2)]^{t_2} \dots [P(C_9)]^{t_9}. \end{aligned} \quad (2)$$

Therefore,

$$\begin{aligned} P\{x < X_{r:n} \leq x + \delta_x, y < Y_{s:n} \leq y + \delta_y\} \\ \equiv P\{\zeta_1 + \zeta_2 + \zeta_3 = r - 1, \zeta_4 + \zeta_5 + \zeta_6 = 1, \\ \zeta_1 + \zeta_4 + \zeta_7 = s - 1, \zeta_2 + \zeta_5 + \zeta_8 = 1\} \\ = \sum_{\substack{t_1+t_2+t_3=r-1 \\ t_4+t_5+t_6=1 \\ t_1+t_4+t_7=s-1 \\ t_2+t_5+t_8=1}} P\{\zeta_1 = t_1, \zeta_2 = t_2, \dots, \zeta_9 = t_9\}. \end{aligned} \quad (3)$$

Let us consider the summation $t_1 + t_2 + t_3 = r - 1$, $t_4 + t_5 + t_6 = 1$, $t_1 + t_4 + t_7 = s - 1$ and $t_2 + t_5 + t_8 = 1$. Because $t_4 + t_5 + t_6 = 1$, this equality is valid if and only if t_4, t_5 and t_6 are 0 or 1. Similarly, t_2, t_5 and t_8 are 0 or 1, because $t_2 + t_5 + t_8 = 1$. Thus, all possible values of (t_4, t_5, t_6) and (t_2, t_5, t_8) are described in the following table:

t_4	t_5	t_6		t_2	t_5	t_8
1	0	0		1	0	0
0	1	0		0	1	0
0	0	1		0	0	1

Because the number of occurrences of C_5 is $t_5 = 1$, from the table, we have $t_4 = t_6 = 0$, $t_2 = t_8 = 0$. If $t_5 = 0$, $t_4 + t_6 = 1$ and $t_2 + t_8 = 1$, where $(t_4 = 1, t_6 = 0)$ or $(t_4 = 0, t_6 = 1)$ and $(t_2 = 1, t_8 = 0)$ or $(t_2 = 0, t_8 = 1)$. Therefore, this summation can be reduced to a simpler form according to cases of $t_5 = 1$ or $t_5 = 0$:

$$\begin{aligned}
& P\{x < X_{r:n} \leq x + \delta_x, y < Y_{s:n} \leq y + \delta_y\} \\
&= \sum_{\substack{t_1+t_2+t_3=r-1 \\ t_4+t_5+t_6=1 \\ t_1+t_4+t_7=s-1 \\ t_2+t_5+t_8=1}} P\{\zeta_1 = t_1, \zeta_2 = t_2, \dots, \zeta_9 = t_9\} \\
&= \sum_{t_1=a_1}^{a_2} P(t_1; r, s, n) + \sum_{t_4=d_1}^{d_2} \sum_{t_2=c_1}^{c_2} \sum_{t_1=b_1}^{b_2} P(t_1, t_2, t_4; r, s, n), \quad (4)
\end{aligned}$$

where

$$\begin{aligned}
& P(t_1; r, s, n) \\
&\equiv P\{\zeta_1 = t_1, \zeta_2 = 0, \zeta_3 = r - 1 - t_1, \zeta_4 = 0, \zeta_5 = 1, \\
&\quad \zeta_6 = 0, \zeta_7 = s - 1 - t_1, \zeta_8 = 0, \zeta_9 = n - r - s + t_1 + 1\};
\end{aligned}$$

and

$$\begin{aligned}
& P(t_1, t_2, t_4; r, s, n) \\
&\equiv P\{\zeta_1 = t_1, \zeta_2 = t_2, \zeta_3 = r - 1 - t_1 - t_2, \zeta_4 = t_4, \zeta_5 = 0, \\
&\quad \zeta_6 = 1 - t_4, \zeta_7 = s - 1 - t_1 - t_4, \zeta_8 = 1 - t_2, \\
&\quad \zeta_9 = n - r - s + t_1 + t_2 + t_4\}.
\end{aligned}$$

Therefore, by using (2) we obtain

$$\begin{aligned}
& P(t_1; r, s, n) \\
&= \frac{n!}{t_1!(r-1-t_1)!(s-1-t_1)!(n-r-s+t_1+1)!} \\
&\quad \times [F(x, y)]^{t_1} [F_X(x) - F(x, y + \delta_y)]^{r-1-t_1} \\
&\quad \times [F(x + \delta_x, y + \delta_y) - F(x + \delta_x, y) - F(x, y + \delta_y) + F(x, y)] \\
&\quad \times [F_Y(y) - F(x + \delta_x, y)]^{s-1-t_1} \\
&\quad \times [1 - F_X(x + \delta_x) - F_Y(y + \delta_y) + F(x + \delta_x, y + \delta_y)]^{n-r-s+t_1+1} \quad (5)
\end{aligned}$$

and

$$\begin{aligned}
& P(t_1, t_2, t_4; r, s, n) \\
= & \frac{n!}{t_1!(r-1-t_1-t_2)!(s-1-t_1-t_4)!(n-r-s+t_1+t_2+t_4)!} \\
& \times [F(x, y)]^{t_1} [F(x, y + \delta_y) - F(x, y)]^{t_2} \\
& \times [F_X(x) - F(x, y + \delta_y)]^{r-1-t_1-t_2} [F(x + \delta_x, y) - F(x, y)]^{t_4} \\
& \times [F_X(x + \delta_x) - F_X(x) - F(x + \delta_x, y + \delta_y) + F(x, y + \delta_y)]^{1-t_4} \\
& \times [F_Y(y) - F(x + \delta_x, y)]^{s-1-t_1-t_4} \\
& \times [F_Y(y + \delta_y) - F_Y(y) - F(x + \delta_x, y + \delta_y) + F(x + \delta_x, y)]^{1-t_2} \\
& \times [1 - F_X(x + \delta_x) - F_Y(y + \delta_y) + F(x + \delta_x, y + \delta_y)]^{n-r-s+t_1+t_2+t_4}.
\end{aligned} \tag{6}$$

Finally, by substituting the (5) and (6) in the equation (4) and taking the limit as $\delta_x, \delta_y \rightarrow 0$, we obtain the joint density function of bivariate order statistics $X_{r:n}$ and $Y_{s:n}$.